

# Large deviations in random graphs

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(joint work with Matan Harel, Gady Kozma, and Frank Mousset)

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### Question

How 'concentrated' is  $X_N$  around its expectation?

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Two possible ways to strengthen this result:

- How fast can  $\varepsilon$  tend to zero as  $N \rightarrow \infty$ ?
- What is the rate of convergence?

## Typical deviations – Central Limit Theorem

The standard deviation  $\sigma$  of  $Y_1$  is defined by

$$\sigma := \sqrt{\text{Var}(Y_1)} = (\mathbb{E}[Y_1^2] - \mathbb{E}[Y_1]^2)^{1/2}.$$

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It is already unlikely that  $|X_N - \mu N| \gg \sqrt{N}$ .

The limiting behaviour depends only on  $\mathbb{E}[Y_1]$  and  $\mathbb{E}[Y_1^2]$ .

## Large deviations – Cramér's theorem

### Theorem (Cramér 1938)

There is a function  $I = I_{Y_1} : (0, \infty) \rightarrow (0, \infty]$  such that

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as  $Y_1, \dots, Y_N$  are i.i.d. We choose the optimal value of  $\lambda$  (...)



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A natural example coming from random graph theory:

$$X_N = \#\text{triangles in } G_{n,p};$$

here,  $N = \binom{n}{2}$  and  $X_N$  may be expressed as degree-three polynomial in  $N$  independent Bernoulli random variables.

## Triangles in the binomial random graph

The binomial random graph  $G_{n,p}$  has vertex set  $\llbracket n \rrbracket := \{1, \dots, n\}$  and

$$\Pr(ij \in G_{n,p}) = p \quad \text{for all } i, j \in \llbracket n \rrbracket,$$

independently of all other pairs.

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Let  $X_N$  denote the number of triangles in  $G_{n,p}$  and note that

$$X_N = \sum_{i,j,k} Y_{ij} Y_{ik} Y_{jk} \quad \text{and} \quad \mathbb{E}[X_N] = \binom{n}{3} p^3,$$

where  $Y_{ij} = \mathbf{1}_{ij \in G_{n,p}} \sim \text{Bernoulli}(p)$ .

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## Remark

We will allow  $p$  to depend on  $n$ . In fact, assume  $p = p(n) \rightarrow 0$  as  $n \rightarrow \infty$ .



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where  $\sigma_N$  is the standard deviation of  $X_N$ .

The standard deviation of  $X_N$  is straightforward to compute:

$$\sigma_N^2 = \text{Var}(X_N) = \binom{n}{3} p^3 (1-p^3) + \binom{n}{4} \binom{4}{2} p^5 (1-p).$$

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We conclude that

$$\Pr(X \geq (1 + \delta)\mathbb{E}[X]) \geq \exp(-c_\delta n^2 p^2 \log(1/p)).$$

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### Theorem (Chatterjee / DeMarco–Kahn)

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### Theorem (Chatterjee / DeMarco–Kahn)

If  $p \gg \log n/n$ , then, for every fixed  $\delta > 0$ ,

$$\Pr(X \geq (1 + \delta)\mathbb{E}[X]) = \exp(-\Theta_\delta(n^2 p^2 \log(1/p))).$$

The assumption  $p \gg \log n/n$  is necessary.

## Upper tail – lower bounds (revisited)

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## Theorem (Lubetzky–Zhao 2014)

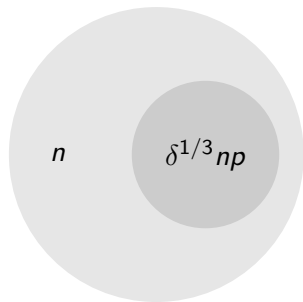
$$\psi(\delta)/n^2 p^2 \rightarrow \begin{cases} \delta^{2/3}/2 & \text{if } n^{-1} \ll p \ll n^{-1/2}, \\ \min\{\delta^{2/3}/2, \delta/3\} & \text{if } n^{-1/2} \ll p \ll 1. \end{cases}$$

## Optimal planted subgraphs

The constants  $\delta^{2/3}/2$  and  $\delta/3$  come from the following:

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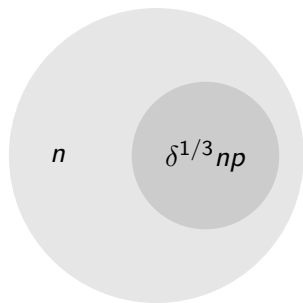
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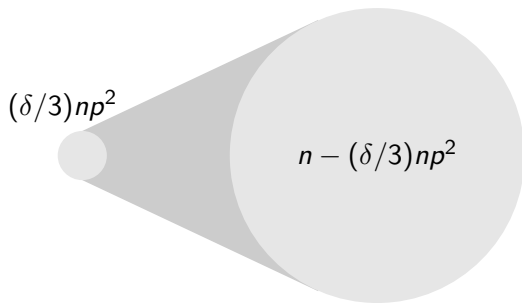
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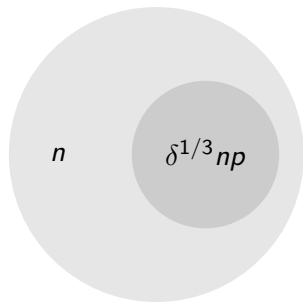
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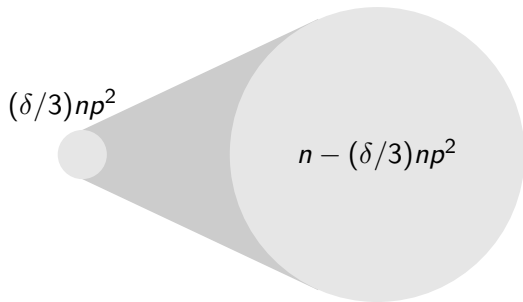
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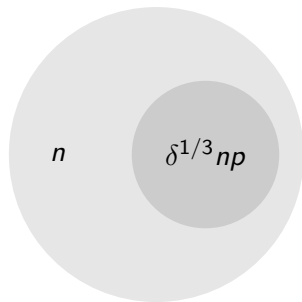
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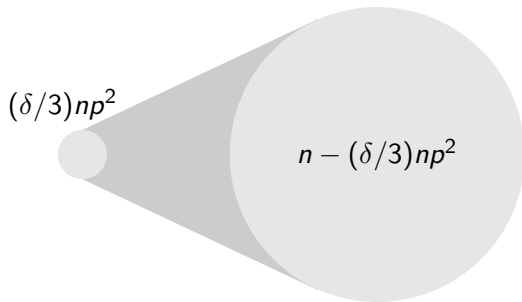
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The 'hub' works only when  $np^2 \gg 1$ , as  $(\delta/3)np^2$  is assumed an integer.



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We expect the following to be true (the assumption  $p \ll 1$  is needed):

### Theorem

If  $n^{-\alpha} \ll p \ll 1$ , then, for every  $\delta > 0$ ,

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where  $\text{Po}(\delta) = (1 + \delta) \log(1 + \delta) - \delta$ .

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### Corollary (Harris 1960 / Janson 1990)

If  $p < .99$ , then, for every  $\delta \in (0, 1]$ ,

$$\Pr(X \leq (1 - \delta)\mathbb{E}[X]) = \exp(-\Theta_\delta(\min\{n^2 p, n^3 p^3\})).$$



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### Theorem (Łuczak 2000)

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If  $\delta < 1$ , then we could consider a graph  $G_\delta$  with at most  $(1 - \delta)\binom{n}{3}$  triangles and as many edges as possible to obtain

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### Proposition

Suppose that  $q$  is such that  $\mathbb{E}[\#K_3(G_{n,q})] \leq (1 - \delta)\mathbb{E}[X] = (1 - \delta)\binom{n}{3}p^3$ .  
Then,

$$\Pr(X \leq (1 - \delta)\mathbb{E}[X]) \geq \exp\left(- (1 + o(1)) \cdot \sum_{i,j} I_p(q_{ij})\right),$$

where  $I_p(q) = q \log \frac{q}{p} + (1 - q) \log \frac{1-q}{1-p}$ .



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Define, for every  $\delta \in (0, 1]$ ,

$$\Phi(\delta) = \min \left\{ \sum_{i,j} I_p(q_{ij}) : \mathbb{E}[\#K_3(G_{n,q})] \leq (1 - \delta)\mathbb{E}[X] \right\}.$$

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**Thank you for your attention!**